

Stat 544 Exam 1

March 1, 2007

**I have neither given nor received unauthorized assistance on this examination.**

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\_\_\_\_\_ date

1. Suppose that a parameter  $\theta = (\theta_1, \theta_2)$  takes one of the values in  $\left\{ (1,0), \left(2, \frac{1}{4}\right), \left(2, \frac{3}{4}\right) \right\}$ . Then  $X$
- is uniformly distributed on  $\{0,1,2\}$  if  $\theta_1 = 1$
  - is Binomial $(2, \theta_2)$  if  $\theta_1 = 2$
- (\*)

A Bayesian uses prior probabilities

$$g((1,0)) = \frac{1}{2}, \quad g\left(\left(2, \frac{1}{4}\right)\right) = \frac{1}{4}, \quad \text{and} \quad g\left(\left(2, \frac{3}{4}\right)\right) = \frac{1}{4}$$

- a) In the table below give the probability mass function of the (joint) distribution of  $(X, \theta)$  that the Bayesian is using. (No need to do arithmetic.)

		$x$		
		0	1	2
$\theta$	$(1,0)$			
	$\left(2, \frac{1}{4}\right)$			
	$\left(2, \frac{3}{4}\right)$			

- b) Give below (in tabular form: possible values and posterior probabilities) the posterior distributions of first  $\theta$  and then  $\theta_1$  if  $X = 2$  (say,  $g(\theta|2)$  and then  $g(\theta_1|2)$ ).

- c) Suppose that given  $\theta$ ,  $X$  and  $X_{\text{new}}$  are independent with marginal distributions specified as in (\*). Find the posterior predictive distribution of  $X_{\text{new}}$  given that  $X = 2$  and give it below in tabular form.

2. Call the  $p$  quantile of the  $\chi_v^2$  distribution  $\chi_{v,p}^2$  and use this notation in your answers below.

a) A Bayesian observes 5 iid  $N(0, \sigma^2)$  observations  $X_1, \dots, X_5$  and finds that  $\sum_{i=1}^5 X_i^2 = 13$ . What are 95% credible limits for  $\sigma^2$  that would be reported by this person if he or she is using the Jeffreys improper prior for  $\sigma^2$ ?

b) A Bayesian observes 5 iid  $MVN_2(\mathbf{0}, \Sigma)$  random vectors  $\mathbf{X}_1, \dots, \mathbf{X}_5$  and finds

$\mathbf{S}_0 = \sum_{i=1}^5 \mathbf{X}_i \mathbf{X}_i' = \begin{pmatrix} 13 & 3 \\ 3 & 10 \end{pmatrix}$ . What are 95% credible limits for  $\sigma_{11}$  the *variance* that is the upper left entry of  $\Sigma$  if he or she is using an inv-Wishart $(3, (.1\mathbf{I}))$  proper prior for  $\Sigma$ ?

3. The two-parameter double exponential distribution is a continuous distribution with pdf on  $\mathfrak{R}$

$$f(x | \mu, \beta) = \frac{\beta}{2} \exp(-\beta|x - \mu|)$$

(this distribution is symmetric about  $\mu$ , and  $\beta$  controls the distribution variance). Suppose that a Bayesian observes  $X_1, X_2, \dots, X_n$  that are iid according to this distribution and decides to use an improper prior for the parameter vector  $(\mu, \beta)$  of the form

$$g(\mu, \beta) \propto 1 \cdot \exp(-\gamma\beta)$$

for some real number  $\gamma$ .

a) Carefully describe a Gibbs algorithm for sampling from the posterior  $g(\mu, \beta | \mathbf{x})$ . (Give the forms of the pdf's from which one must sample for each update. If a simulation from a standard distribution can be employed, say exactly which distribution one may use.)

b) Carefully describe any Metropolis-Hastings algorithm that can be used to sample from the posterior  $g(\mu, \beta | \mathbf{x})$ .

4. An absent-minded gambler plays a particular slot machine daily for about 2 hours per day. This person is interested in the "win" rate for the machine (the probability,  $p$ , that any particular play results in some prize), but only records how many "wins" he sees in a day (and not exactly how many times he plays). Attached to this exam is a WinBUGS printout for two Bayes analyses based on

$$X_i = \text{number of wins on day } i$$

for 10 days of play ( $i = 1, 2, \dots, 10$ ) and consistent with the rough information that physical limitations on the machine and the gambler make it essentially certain that each

$$n_i = \text{the number of plays made on day } i$$

was somewhere between 150 and 250. (These are for Models **A** and **B**.)

a) For these two analyses, carefully (and fully) state the model assumptions used:

	Model A	Model B
(conditional) distribution of $(X_1, \dots, X_{10})   (p, n_1, \dots, n_{10})$		
(prior) distribution of $p$		
(prior) distribution of $(n_1, \dots, n_{10})$		
relation of the priors for $p$ and $(n_1, \dots, n_{10})$		

b) In what specific way(s) is the difference in the prior assumptions on  $(n_1, \dots, n_{10})$  reflected in the posteriors for the  $n_i$ ?

c) How (if at all) is the difference in the prior assumptions on  $(n_1, \dots, n_{10})$  reflected in the posteriors for  $p$ ? Why is this reasonable?

d) The gambler likes the Model A analysis. What is an interval that he believes is 95% sure to contain the number of wins in a standard 2 hour session with the machine *tomorrow*?

e) A substantially different analysis was run using what we might call Model C, with WinBUGS code and some summaries below. Compare these results to the Model A and Model B results and discuss why both the differences and the similarities you see are (in retrospect) reasonable.

**Model C**

```

model {
lambda~dexp(.005)
for (i in 1:11) {
  n[i]~dpois(lambda)
}
p~dunif(0,1)
for (i in 1:11) {
  X[i]~dbin(p,n[i])
}
}
list(X=c(9,10,14,10,11,14,11,7,11,17,NA))

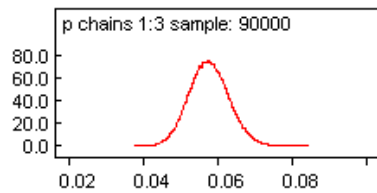
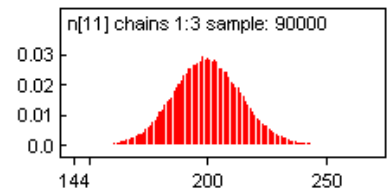
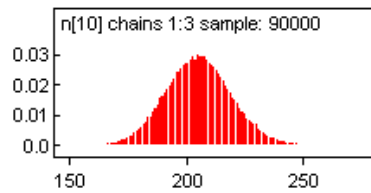
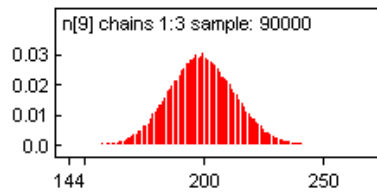
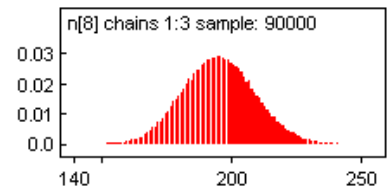
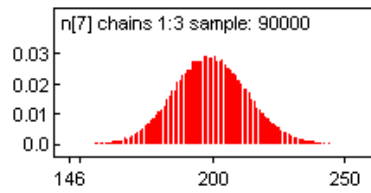
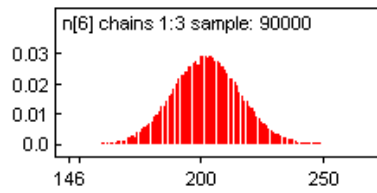
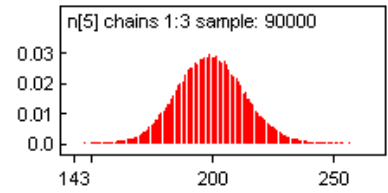
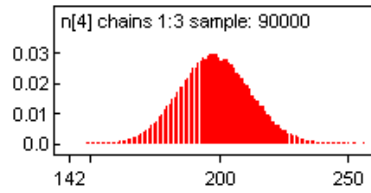
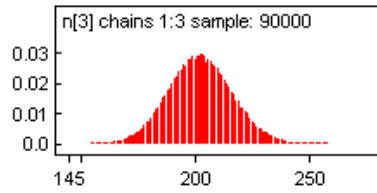
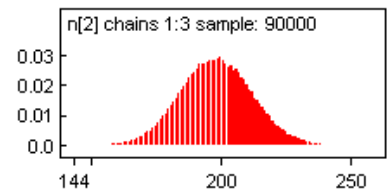
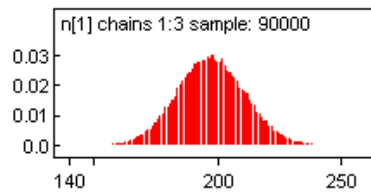
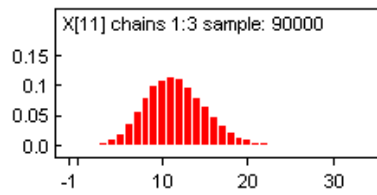
```

node	mean	sd	MC error	2.5%	median	97.5%	start	sample
X[11]	11.46	3.551	0.00377	5.0	11.0	19.0	61001	900000
lambda	77.66	94.39	2.194	11.91	42.12	341.7	61001	900000
n[1]	75.21	94.73	2.195	9.0	40.0	340.0	61001	900000
n[2]	76.19	94.7	2.194	10.0	41.0	341.0	61001	900000
⋮								
n[10]	83.19	94.7	2.194	17.0	48.0	348.0	61001	900000
n[11]	77.66	94.79	2.194	10.0	42.0	343.0	61001	900000
p	0.3493	0.2683	0.005312	0.03306	0.2715	0.942	61001	900000

## Model A

```
model {  
  
  for (i in 1:11) {  
    n[i]~dpois(200)  
  }  
  
  p~dbeta(1,1)  
  
  for (i in 1:11) {  
    X[i]~dbin(p,n[i])  
  }  
  
}  
  
list(X=c(9,10,14,10,11,14,11,7,11,17,NA))
```

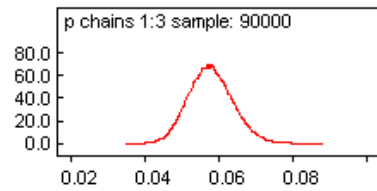
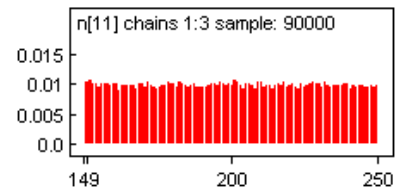
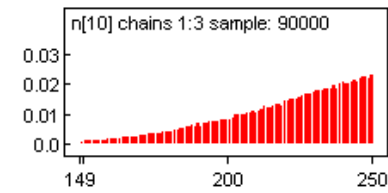
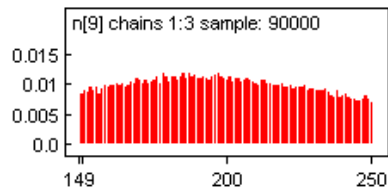
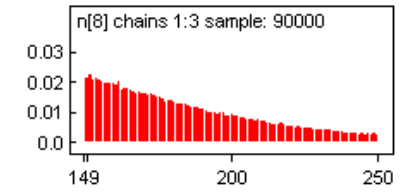
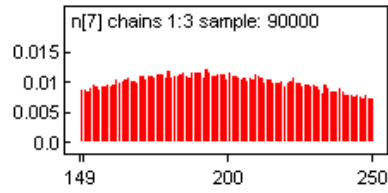
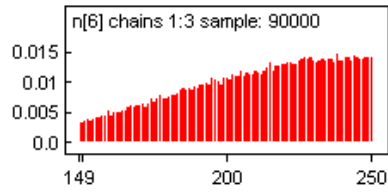
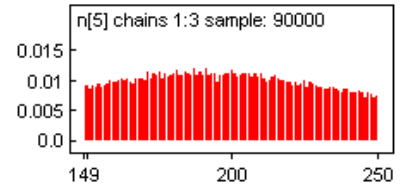
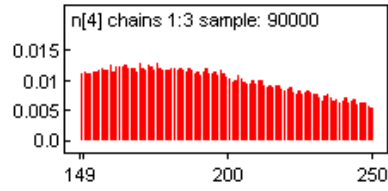
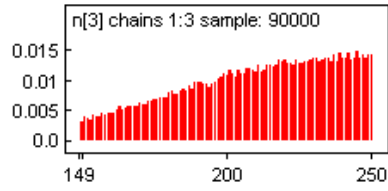
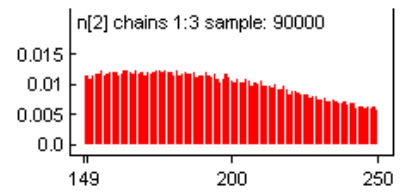
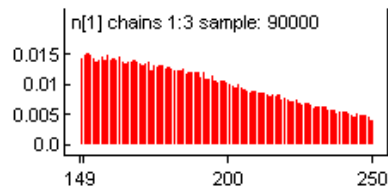
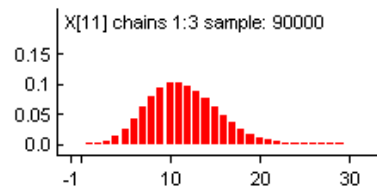
<b>node</b>	<b>mean</b>	<b>sd</b>	<b>MC error</b>	<b>2.5%</b>	<b>median</b>	<b>97.5%</b>	<b>start</b>	<b>sample</b>
X[11]	11.51	3.545	0.01181	5.0	11.0	19.0	31001	90000
n[1]	197.4	13.78	0.04966	171.0	197.0	225.0	31001	90000
n[2]	198.5	13.8	0.04444	172.0	198.0	226.0	31001	90000
n[3]	202.5	13.75	0.04545	176.0	202.0	230.0	31001	90000
n[4]	198.5	13.7	0.04662	172.0	198.0	226.0	31001	90000
n[5]	199.4	13.8	0.04601	173.0	199.0	227.0	31001	90000
n[6]	202.5	13.83	0.04462	176.0	202.0	230.0	31001	90000
n[7]	199.5	13.77	0.04515	173.0	199.0	227.0	31001	90000
n[8]	195.5	13.74	0.04557	169.0	195.0	223.0	31001	90000
n[9]	199.5	13.76	0.04263	173.0	199.0	227.0	31001	90000
n[10]	205.5	13.75	0.04313	179.0	205.0	233.0	31001	90000
n[11]	199.9	14.11	0.04914	173.0	200.0	228.0	31001	90000
p	0.05755	0.00535	1.953E-5	0.04747	0.0574	0.06847	31001	90000



## Model B

```
model {  
  
  for (j in 1:101) {  
    alpha[j] <- (1/101)  
  }  
  
  for (i in 1:11) {  
    m[i] ~ dcat(alpha[])  
    n[i] <- m[i] + 149  
  }  
  
  p ~ dunif(0,1)  
  
  for (i in 1:11) {  
    X[i] ~ dbin(p, n[i])  
  }  
  
}  
  
list(X=c(9,10,14,10,11,14,11,7,11,17,NA))
```

<b>node</b>	<b>mean</b>	<b>sd</b>	<b>MC error</b>	<b>2.5%</b>	<b>median</b>	<b>97.5%</b>	<b>start</b>	<b>sample</b>
X[11]	11.54	3.888	0.01362	5.0	11.0	20.0	31001	90000
n[1]	190.3	27.21	0.09146	151.0	187.0	245.0	31001	90000
n[2]	194.4	27.66	0.09497	152.0	192.0	246.0	31001	90000
n[3]	210.0	26.54	0.0897	156.0	213.0	249.0	31001	90000
n[4]	194.2	27.67	0.09264	152.0	192.0	246.0	31001	90000
n[5]	198.2	27.87	0.09913	152.0	197.0	247.0	31001	90000
n[6]	210.0	26.54	0.09091	156.0	214.0	249.0	31001	90000
n[7]	198.4	27.79	0.09596	152.0	198.0	247.0	31001	90000
n[8]	183.1	25.41	0.08395	151.0	177.0	241.0	31001	90000
n[9]	198.3	27.81	0.09577	152.0	197.0	247.0	31001	90000
n[10]	219.6	23.33	0.08045	165.0	225.0	249.0	31001	90000
n[11]	199.9	29.15	0.09933	152.0	200.0	248.0	31001	90000
p	0.05768	0.005921	2.393E-5	0.04678	0.05747	0.06995	31001	90000



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