

STAT 511 HW#6 SPRING 2009

PROBLEM 1:

```
y<-c(6.0,6.1,8.6,7.1,6.5,7.4,9.4,9.9,9.5,7.5,6.4,9.1,8.7)
A<-c(1,1,1,1,1,1,2,2,2,2,2,2,2)
B<-c(1,1,2,2,2,2,1,1,2,2,3,3,3)
options(contrasts=c("contr.sum","contr.sum"))
```

a)

The pooled variance from the 5 samples (of sizes 2,4,2,2, and 3) is 1.089(standard error =1.044). An exact 95% confidence interval for σ is

$$\left(\sqrt{\frac{1.089*8}{\chi_{0.975,60}^2}}, \sqrt{\frac{1.089*8}{\chi_{0.025,60}^2}} \right) = (0.705, 2)$$

##Using lme

```
A<-as.factor(A)
```

```
B<-as.factor(B)
```

```
lme.out<-lme(y~1,random=~1|A/B)
```

```
summary(lme.out)
```

Linear mixed-effects model fit by REML

Data: NULL

AIC	BIC	logLik
49.46496	51.40458	-20.73248

Random effects:

Formula: ~1 | A
(Intercept)

StdDev: 1.165728

Formula: ~1 | B %in% A
(Intercept) Residual

StdDev: 0.4955114 1.056588

Fixed effects: y ~ 1

	Value	Std.Error	DF	t-value	p-value
(Intercept)	7.79261	0.9051327	8	8.609356	0

Standardized Within-Group Residuals:

Min	Q1	Med	Q3	Max
-1.8490691	-0.6716102	0.1801882	0.7063261	1.3159194

Number of Observations: 13

Number of Groups:

A	B	%in% A
2	5	

```
intervals(lme.out)
```

Approximate 95% confidence intervals

Fixed effects:

	lower	est.	upper
(Intercept)	5.70537	7.79261	9.87985

Random Effects:

Level: A

	lower	est.	upper
sd((Intercept))	0.2184694	1.165728	6.220191

Level: B

	lower	est.	upper
sd((Intercept))	0.03991159	0.4955114	6.151885

Within-group standard error:

	lower	est.	upper
	0.6391959	1.0565880	1.7465355

An approximate 95% confidence interval for σ is (0.6391959 , 1.7465355)

```
##Using lmer
lmer.out<- lmer(y ~ (1|A) + (1|A:B))
summary(lmer.out)
Linear mixed model fit by REML
Formula: y ~ (1 | A) + (1 | A:B)
  AIC   BIC logLik deviance REMLdev
 49.46 51.72 -20.73   43.06   41.46
Random effects:
  Groups   Name      Variance Std.Dev.
  A:B      (Intercept) 0.24553  0.49551
  A        (Intercept) 1.35892  1.16573
  Residual                    1.11638  1.05659
Number of obs: 13, groups: A:B, 5; A, 2
```

```
Fixed effects:
              Estimate Std. Error t value
(Intercept)   7.7926     0.9051     8.61
sims <- mcmcscamp(lmer.out, 50000)
HPDinterval(sims)
$fixef
              lower      upper
(Intercept) 5.006724 10.53065
$ST
      lower      upper
[1,]    0 0.992304
[2,]    0 3.311172
$sigma
      lower      upper
[1,] 0.7268979 1.814918
```

b)

```
AA<-as.factor(A)
BB<-as.factor(B)
lm.out<-lm(y~1+AA/BB)
summary(lm.out)
Call:
lm(formula = y ~ 1 + AA/BB)

Residuals:
      Min       1Q   Median       3Q      Max
-1.667e+00 -3.000e-01  4.612e-16  6.333e-01  1.200e+00

Coefficients: (1 not defined because of singularities)
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   8.0694     0.3293  24.503 8.22e-09 ***
AA1          -0.6694     0.3293  -2.033  0.0765 .
AA1:BB1      -1.3500     0.9040  -1.493  0.1737
AA2:BB1       0.9111     0.5857   1.556  0.1584
AA1:BB2       NA          NA       NA     NA
AA2:BB2      -0.2389     0.5857  -0.408  0.6941
Residual standard error: 1.044 on 8 degrees of freedom
Multiple R-squared:  0.6286,    Adjusted R-squared:  0.4429
F-statistic: 3.385 on 4 and 8 DF,  p-value: 0.06686
```

Note that the estimate of σ (residual standard error = 1.044) produced by lm.out is exactly the one based on the mean square error referred in part a).

c)

The predictions for the fixed effects model correspond to the cell sample means \bar{y}_{ij} .

```

round(predict(lm.out), 3)
  1      2      3      4      5      6      7      8      9     10     11     12     13
6.050 6.050 7.400 7.400 7.400 7.400 9.650 9.650 8.500 8.500 8.067 8.067 8.067
round(predict(lme.out), 3)
  1/1  1/1  1/2  1/2  1/2  1/2  2/1  2/1  2/2  2/2  2/3  2/3  2/3
6.739 6.739 7.210 7.210 7.210 7.210 8.881 8.881 8.530 8.530 8.354 8.354 8.354
random.effects(lme.out)
  Level: A
    (Intercept)
  1 -0.7504841
  2  0.7504841
  Level: B %in% A
    (Intercept)
  1/1 -0.30308804
  1/2  0.16748966
  2/1  0.33815279
  2/2 -0.01316488
  2/3 -0.18938953
round(fitted(lmer.out), 3)
7.446 7.446 7.609 7.609 7.609 7.609 8.632 8.632 8.664 8.664 7.814 7.814 7.814
ranef(lmer.out)
$`A:B`
  (Intercept)
  1:1 -0.17374831
  1:2  0.01033243
  2:1 -0.16270189
  2:2 -0.12574233
  2:3 -1.08660931
$A
  (Intercept)
  1 -0.9928770
  2  0.3693292

```

PROBLEM 2

The weights of the (75) pieces of silver were determined and a standard hierarchical (balanced data) ANOVA table was produced as below. (The units of weight were not given in the original article.)

Term	Source	SS	df	E (MS)
A	Machines	1966	2	$\sigma^2 + 5\sigma_\beta^2 + 25\sigma_\alpha^2$
B(A)	Rolls	644	12	$\sigma^2 + 5\sigma_\beta^2$
Error	Pieces	280	60	σ^2
	Total	2890	74	

a)

- $\hat{\sigma}^2 = MSE = 4.667$

A 95% confidence interval for σ is $\left(\sqrt{\frac{SSE}{\chi_{0.975,60}^2}}, \sqrt{\frac{SSE}{\chi_{0.025,60}^2}} \right) = (1.833, 2.630)$

- $\hat{\sigma}_\beta^2 = \frac{MSB(A) - MSE}{5} = 9.8$

$$\hat{v} = \frac{(\hat{\sigma}_\beta^2)^2}{\frac{(MSB(A)/5)^2}{12} + \frac{(MSE/5)^2}{60}} = 9.989$$

A 95% confidence interval for σ_β is $\left(\sqrt{\frac{\hat{v}\hat{\sigma}_\beta^2}{\chi_{0.975,\hat{v}}^2}}, \sqrt{\frac{\hat{v}\hat{\sigma}_\beta^2}{\chi_{0.025,\hat{v}}^2}} \right) = (2.186, 5.501)$

- $\hat{\sigma}_\alpha^2 = \frac{MSA - MSB(A)}{5 \times 5} = 37.173$
 $\hat{v} = \frac{(\hat{\sigma}_\alpha^2)^2}{\frac{(MSA/25)^2}{2} + \frac{(MSB(A)/25)^2}{12}} = 1.787$

A 95% confidence interval for σ_α is $\left(\sqrt{\frac{\hat{v}\hat{\sigma}_\alpha^2}{\chi_{0.975,\hat{v}}^2}}, \sqrt{\frac{\hat{v}\hat{\sigma}_\alpha^2}{\chi_{0.025,\hat{v}}^2}} \right) = (3.098, 46.298)$

- The largest part of the variation comes from the difference between machines

b) A 95% confidence interval for μ is $\bar{y}_{..} \pm t_{0.975,2} \sqrt{\frac{MSA}{3 \times 5 \times 5}} = (19.42, 50.58)$

PROBLEM 3

```

Y<-c(12,13,14,20,8,10,6,7,8,10,11,13,7)
a<-c(1,1,1,1,2,2,2,2,2,3,3,3,3)
b<-c(1,2,2,3,1,2,3,3,3,1,1,2,3)
group<-rep(1,length(Y))
A<-as.factor(a)
B<-as.factor(b)
Group<-as.factor(group)
Fake<-data.frame(Y,A,B,Group)
FakeGrouped<-groupedData(Y~1|group,Fake)
##Using lme
twoway.out<-lme(Y~1,data=FakeGrouped,random=pdBlocked(list(pdIdent(~A-1),pdIdent(~B-1))))
summary(twoway.out)
Linear mixed-effects model fit by REML
Data: FakeGrouped
      AIC      BIC    logLik
 71.49413 73.43375 -31.74706
Random effects:
Composite Structure: Blocked
Block 1: A1, A2, A3
Formula: ~A - 1 | group
Structure: Multiple of an Identity
           A1      A2      A3
StdDev: 3.304856 3.304856 3.304856
Block 2: B1, B2, B3
Formula: ~B - 1 | group
Structure: Multiple of an Identity
           B1      B2      B3 Residual
StdDev: 7.969662e-05 7.969662e-05 7.969662e-05 2.574438
Fixed effects: Y ~ 1
              Value Std.Error DF  t-value p-value
(Intercept) 10.90533   2.03851 12  5.349656  2e-04
Standardized Within-Group Residuals:
           Min      Q1      Med      Q3      Max
-1.29594152 -0.48304590 -0.09461158  0.25779573  2.23599429
Number of Observations: 13
Number of Groups: 1

```

```
##Using lmer
twoway.out2<-lmer(Y ~ (1 | A) + (1| B))
summary(twoway.out2)
  Linear mixed model fit by REML
  Formula: Y ~ (1 | A) + (1 | B)
    AIC   BIC logLik deviance REMLdev
  71.5 73.75 -31.75   66.72   63.5
Random effects:
  Groups   Name      Variance  Std.Dev.
  A        (Intercept) 1.0922e+01 3.3049e+00
  B        (Intercept) 4.3566e-20 2.0872e-10
  Residual                    6.6277e+00 2.5744e+00
Number of obs: 13, groups: A, 3; B, 3

  Fixed effects:
              Estimate Std. Error t value
(Intercept)   10.905      2.039     5.35
sims<-mcmcscamp(twoway.out2,50000)
HPDinterval(sims,prob=.95)
  $fixef
              lower      upper
(Intercept) 6.623834 15.33435
  $ST
      lower      upper
[1,]      0 1.812893
[2,]      0 1.235887
  $sigma
      lower      upper
[1,] 1.793329 4.805657
```

PROBLEM 4

```
a)
Y <-
c(13.5,14.8,10.5,11.7,12.9,12.0,8.8,13.5,12.4,16.0,11.3,12.0,14.0,12.5,13.0,13.1
,12.6,12.7,11.0,10.6)
a <- c(rep(1,10),rep(2,10))
b <- c(rep(c(1,1,2,2,3,3,4,4,5,5),2))
A <- as.factor(a)
B <- as.factor(b)
fitlm.out <- lm(Y~A*B)
anova(fitlm.out)
  Analysis of Variance Table

  Response: Y
    Df Sum Sq Mean Sq F value Pr(>F)
  A     1  0.5445   0.5445   0.2598 0.62129
  B     4  2.6920   0.6730   0.3212 0.85751
  A:B   4 24.4980   6.1245   2.9227 0.07695 .
  Residuals 10 20.9550   2.0955
```

The sum squares presented by Vardeman and Jobe are correct.

b) A 95% confidence interval for σ is $\left(\sqrt{\frac{SSE}{\chi_{0.975,10}^2}}, \sqrt{\frac{SSE}{\chi_{0.025,10}^2}}\right) = (1.011, 2.540)$

$$c) \hat{\sigma}_\beta^2 + \hat{\sigma}_{\alpha\beta}^2 = \frac{MSB+MSAB-2MSE}{4} = 0.652$$

$$\hat{v} = \frac{(\hat{\sigma}_\beta^2 + \hat{\sigma}_{\alpha\beta}^2)^2}{\frac{(MSB/4)^2}{4} + \frac{(MSAB/4)^2}{4} + \frac{(MSE/2)^2}{10}} = 0.604$$

A 95% confidence interval for $\sqrt{\hat{\sigma}_\beta^2 + \hat{\sigma}_{\alpha\beta}^2}$ is $\left(\sqrt{\frac{\hat{v}(\hat{\sigma}_\beta^2 + \hat{\sigma}_{\alpha\beta}^2)}{\chi_{0.975, \hat{v}}^2}}, \sqrt{\frac{\hat{v}(\hat{\sigma}_\beta^2 + \hat{\sigma}_{\alpha\beta}^2)}{\chi_{0.025, \hat{v}}^2}} \right) = (0.321, 238.470)$

```
d)
##Using lme
ab<-c(1,1,2,2,3,3,4,4,5,5,6,6,7,7,8,8,9,9,10,10)
group<-c(rep(1,20))
AB<-as.factor(ab)
Group<-as.factor(group)
Gauge<-data.frame(Y,A,B,AB,Group)
GaugeGrouped<-groupedData(Y~1|Group,Gauge)
fitlme.out<-
lme(Y~1,data=GaugeGrouped,random=list(Group=pdBlocked(list(pdIdent(~A-1),pdIdent(~B-1),pdIdent(~AB-1))))
Linear mixed-effects model fit by REML
  Data: GaugeGrouped
      AIC      BIC    logLik
 84.44239 89.16459 -37.22120
Random effects:
Composite Structure: Blocked
Block 1: A1, A2
Formula: ~A - 1 | Group
Structure: Multiple of an Identity
              A1              A2
StdDev: 0.0001327714 0.0001327714
Block 2: B1, B2, B3, B4, B5
Formula: ~B - 1 | Group
Structure: Multiple of an Identity
              B1              B2              B3              B4              B5
StdDev: 9.246162e-08 9.246162e-08 9.246162e-08 9.246162e-08 9.246162e-08
Block 3: AB1, AB2, AB3, AB4, AB5, AB6, AB7, AB8, AB9, AB10
Formula: ~AB - 1 | Group
Structure: Multiple of an Identity
              AB1              AB2              AB3              AB4              AB5              AB6
AB7              AB8              AB9              AB10 Residual
StdDev: 0.7021792 0.7021792 0.7021792 0.7021792 0.7021792 0.7021792
0.7021792 0.7021792 0.7021792 0.7021792 1.447584
Fixed effects: Y ~ 1
              Value Std.Error DF t-value p-value
(Intercept) 12.445  0.392531 19 31.7045      0
Standardized Within-Group Residuals:
              Min              Q1              Med              Q3              Max
-2.23171950 -0.46808911 -0.03495508  0.32702849  2.06786074
Number of Observations: 20
Number of Groups: 1

##Using lmer
fitlmer.out<-lmer(Y ~ (1|A) + (1| B) + (1| A:B))
Linear mixed model fit by REML
Formula: Y ~ (1 | A) + (1 | B) + (1 | A:B)
```

```

AIC   BIC logLik deviance REMLdev
84.44 89.42 -37.22   74.38   74.44
Random effects:
Groups   Name             Variance Std.Dev.
A:B      (Intercept) 0.49306 0.70218
B        (Intercept) 0.00000 0.00000
A        (Intercept) 0.00000 0.00000
Residual                    2.09550 1.44758
Number of obs: 20, groups: A:B, 10; B, 5; A, 2
Fixed effects:
              Estimate Std. Error t value
(Intercept) 12.4450     0.3925   31.7

```

The point estimate of the repeatability standard deviation σ is 1.447. The point estimate of the reproducibility standard deviation $\sqrt{\hat{\sigma}_\beta^2 + \hat{\sigma}_{\alpha\beta}^2} = 0.702$ based on REML estimates. Note that the confidence interval for σ_β is extremely wide. Then, the point estimate of reproducibility standard deviation is not reliable. Our approximate interval from (c) does not do justice to the uncertainty here.

PROBLEM 5

```

y<-c(9.5,11.3,8.4,8.5,6.7,9.4,6.1,12,12.2,13.1,10.7,8.9,10.8,7.4,8.6)
A<-c(rep(1,4),rep(2,3),rep(1,4),rep(2,4))
B<-c(1,2,3,3,1,2,3,1,1,2,3,1,2,3,3)
Block<-c(rep(1,7),rep(2,8))
A<-as.factor(A)
B<-as.factor(B)
FBlock<-as.factor(Block)
options(contrasts=c("contr.sum","contr.sum"))
lmf.out<-lm(y~A*B+FBlock)
summary(lmf.out)

```

```

Residuals:
    Min       1Q   Median       3Q      Max
-0.65238 -0.12143  0.07143  0.12857  0.54762
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   9.5929     0.1040  92.245 2.13e-13 ***
A1             1.2849     0.1040  12.356 1.72e-06 ***
B1            -0.2476     0.1473  -1.681  0.131
B2             1.5571     0.1541  10.104 7.86e-06 ***
FBlock1       -1.0286     0.1053  -9.768 1.01e-05 ***
A1:B1          0.2603     0.1487   1.751  0.118
A1:B2         -0.2349     0.1541  -1.524  0.166
Residual standard error: 0.394 on 8 degrees of freedom
Multiple R-squared:  0.9793,    Adjusted R-squared:  0.9638
F-statistic: 63.07 on 6 and 8 DF,  p-value: 2.665e-06

```

An estimate of σ is the residual standard error (= 0.394).

```

b)
##Using lme
lmef.out<-lme(y~A*B,random=~1|Block)
summary(lmef.out)
Linear mixed-effects model fit by REML
Data: NULL
              AIC          BIC      logLik
43.26254 44.84034 -13.63127

```

```

Random effects:
  Formula: ~1 | Block
            (Intercept) Residual
StdDev:    1.446977 0.3940026
Fixed effects: y ~ A * B
            Value Std.Error DF   t-value p-value
(Intercept)  9.593456 1.0284387  8   9.328175  0.0000
A1            1.284322 0.1039918  8  12.350226  0.0000
B1           -0.246421 0.1472967  8  -1.672959  0.1329
B2            1.556544 0.1541129  8  10.100022  0.0000
A1:B1         0.262713 0.1486697  8   1.767092  0.1152
A1:B2        -0.234322 0.1541129  8  -1.520455  0.1669
intervals(lmerf.out)
Approximate 95% confidence intervals
Fixed effects:
            lower      est.      upper
(Intercept)  7.22187218  9.5934561 11.96503993
A1            1.04451629  1.2843217  1.52412716
B1           -0.58608791 -0.2464212  0.09324546
B2            1.20115891  1.5565439  1.91192898
A1:B1        -0.08011985  0.2627131  0.60554606
A1:B2        -0.58970676 -0.2343217  0.12106331
Random Effects:
  Level: Block
            lower      est.      upper
sd((Intercept)) 0.3566116 1.446977 5.871215
Within-group standard error:
            lower      est.      upper
0.2413789 0.3940026 0.6431303
##Using lmer
lmerf.out <-lmer(y ~ A*B+(1 |FBlock))
summary(lmerf.out)
Linear mixed model fit by REML
Formula: y ~ A * B + (1 | Block)
      AIC   BIC logLik deviance REMLdev
 43.26 48.93 -13.63   16.21   27.26
Random effects:
 Groups   Name      Variance Std.Dev.
Block    (Intercept) 2.09374  1.4470
Residual                0.15524  0.3940
Number of obs: 15, groups: Block, 2
Fixed effects:
            Estimate Std. Error t value
(Intercept)  9.5935     1.0284   9.329
A1            1.2843     0.1040  12.350
B1           -0.2464     0.1473  -1.673
B2            1.5565     0.1541  10.100
A1:B1         0.2627     0.1487   1.767
A1:B2        -0.2343     0.1541  -1.520
Correlation of Fixed Effects:
      (Intr) A1      B1      B2      A1:B1
A1    -0.007
B1     0.000 -0.098
B2     0.013  0.047 -0.573
A1:B1 -0.009 -0.009 -0.119  0.057
A1:B2  0.005  0.132  0.066 -0.032 -0.559

```


The inferences for the fixed effects are the same in both analyses. They produce the same point estimates for the fixed effects and their standard errors. The point estimates for σ are the same in both analyses, but from lmf.out fitted

model we obtained $\left(\frac{0.394\sqrt{8}}{\sqrt{\chi_{0.975,8}^2}}, \frac{0.394\sqrt{8}}{\sqrt{\chi_{0.025,8}^2}} \right) = (0.266, 0.755)$ as a 95% confidence interval

for σ , and for lmef.out fitted model we obtained (0.241, 0.643), a narrower interval.

In the fixed effects analysis the estimate of the block effect is -1.0286. In the mixed effect model σ_τ is assumed normally distributed with mean zero and the estimate of σ_τ is 1.447. From this distribution, the value -1.0286 is within one standard deviation from its mean and, therefore consistent with the estimate of σ_τ .

c)

random.effects(lmef.out)

(Intercept)

1 -1.017791

2 1.017791

The BLUE of $\mu + \alpha_1 + \beta_1 + \alpha\beta_{11} + \tau_1$ is $9.592857 + 1.284920 - 0.2476190 + 0.2603175 - 1.028571 = 9.862$.

The BLUP of $\mu + \alpha_1 + \beta_1 + \alpha\beta_{11} + \tau_1$ is $9.593456 + 1.284322 - 0.2464212 + 0.2627131 - 1.017791 = 9.876$.

The BLUP of $\mu + \alpha_1 + \beta_1 + \alpha\beta_{11} + \tau_3$ is $9.593456 + 1.284322 - 0.2464212 + 0.2627131 + 0 = 10.894$.

In the fixed effects model we require information from the third block in order to estimate its effect and to predict a new observation from this block.

PROBLEM 6

```
time<-c(rep(1,10), rep(2,10), rep(3,10))
```

```
ad<-c(rep(c(rep(1,5), rep(2,5)), 3))
```

```
test<-rep(1:5, 6)
```

```
loc<-rep(1:10, 3)
```

```
y<-
```

```
c(958,1005,351,549,730,780,229,883,624,375,1047,1122,436,632,784,897,275,964,695,436,933,986,339,512,707,718,202,817,599,351)
```

a)

```
TIME<-as.factor(time)
```

```
AD<-as.factor(ad)
```

```
TEST<-as.factor(test)
```

```
LOC<-as.factor(loc)
```

```
lmf1.out<-lm(y~1+TIME*(AD/TEST))
```

```
anova(lmf1.out)
```

```
Analysis of Variance Table
```

```
Response: y
```

	Df	Sum Sq (SS)	Mean Sq (MS)	Class Name
TIME	2	67073	33537	A
AD	1	168151	168151	B
AD:TEST	8	1833681	229210	C(B)
TIME:AD	2	391	196	AB
TIME:AD:TEST	16	5727	358	Error
Residuals	0	0		

```
lmf2.out<-lm(y~1+TIME*AD*TEST)
```

```
anova(lmf2.out)
```

Analysis of Variance Table

Response: y

	Df	Sum Sq (SS)	Mean Sq (MS)	Class Name
TIME	2	67073	33537	A
AD	1	168151	168151	B
TEST	4	397490	99373	
TIME:AD	2	391	196	AB
TIME:TEST	8	2475	309	
AD:TEST	4	1436191	359048	
TIME:AD:TEST	8	3252	407	
Residuals	0	0		

SSC(B) = SS(TEST) + SS(AD:TEST) and Df(C(B)) = Df(TEST) + Df(AD:TEST).

SSE = SS(TIME:TEST) + SS(TIME:AD:TEST) and Df(Error) = Df(TIME:TEST) + Df(TIME:AD:TEST).

b)

##With lme

```
lmef.out<-lme(y~1+TIME*AD,random=~1|LOC)
```

```
##equivalent to lmef.out<-lme(y~1+TIME*AD,random=~1|AD/TEST)
```

```
summary(lmef.out)
```

Linear mixed-effects model fit by REML

Data: NULL

	AIC	BIC	logLik
	295.1452	304.5696	-139.5726

Random effects:

Formula: ~1 | LOC

(Intercept) Residual

StdDev: 276.1957 18.92001

Fixed effects: y ~ 1 + TIME * AD

	Value	Std.Error	DF	t-value	p-value
(Intercept)	664.5333	87.40902	16	7.602572	0.0000
TIME1	-16.1333	4.88512	16	-3.302543	0.0045
TIME2	64.2667	4.88512	16	13.155583	0.0000
AD1	74.8667	87.40902	8	0.856510	0.4166
TIME1:AD1	-4.6667	4.88512	16	-0.955281	0.3536
TIME2:AD1	0.5333	4.88512	16	0.109175	0.9144

Correlation:

	(Intr)	TIME1	TIME2	AD1	TIME1:
TIME1	0.0				
TIME2	0.0	-0.5			
AD1	0.0	0.0	0.0		
TIME1:AD1	0.0	0.0	0.0	0.0	
TIME2:AD1	0.0	0.0	0.0	0.0	-0.5

Standardized Within-Group Residuals:

	Min	Q1	Med	Q3	Max
	-1.46621417	-0.41184117	-0.03544604	0.50875325	1.86359419

Number of Observations: 30

Number of Groups: 10

```
intervals(lmef.out)
```

Approximate 95% confidence intervals

Fixed effects:

	lower	est.	upper
(Intercept)	479.234496	664.5333333	849.832170
TIME1	-26.489336	-16.1333333	-5.777331
TIME2	53.910664	64.2666667	74.622669
AD1	-126.698886	74.8666667	276.432220

```

TIME1:AD1    -15.022669  -4.6666667   5.689336
TIME2:AD1    -9.822669   0.5333333  10.889336
Random Effects:
  Level: LOC
              lower    est.    upper
sd((Intercept)) 169.0765 276.1957 451.1805
  Within-group standard error:
    lower    est.    upper
13.37976 18.92001 26.75434
predict(lmef.out)
958.1586 1016.4009 355.1019 543.8067 719.5319 786.5408 224.4200
876.0674 627.7891 376.1827 1043.7586 1102.0009 440.7019 629.4067
805.1319 861.7408 299.6200 951.2674 702.9891 451.3827 934.9586
993.2009 331.9019 520.6067 696.3319 745.7408 183.6200 835.2674
586.9891 335.3827

```

An approximate 95% confidence interval for σ_j is (169.076, 451.181) and an approximate 95% confidence interval for σ is (13.380, 26.754).

```

##With lmer
lmerf.out <-lmer(y ~ TIME*AD +(1 |LOC))
##equivalent to lmerf.out <-lmer(y ~ TIME*AD +(1 |TEST:AD))
summary(lmerf.out)
Linear mixed model fit by REML
Formula: y ~ TIME * AD + (1 | LOC)
   AIC   BIC logLik deviance REMLdev
 295.1 306.4 -139.6   319.5   279.1
Random effects:
 Groups   Name      Variance Std.Dev.
 LOC      (Intercept) 76284.10 276.20
 Residual                    357.97 18.92
Number of obs: 30, groups: LOC, 10
Fixed effects:
              Estimate Std. Error t value
(Intercept)  664.5333    87.4038   7.603
TIME1        -16.1333    4.8851  -3.303
TIME2         64.2667    4.8851  13.156
AD1           74.8667    87.4038   0.857
TIME1:AD1    -4.6667    4.8851  -0.955
TIME2:AD1     0.5333    4.8851   0.109

```

c)

- $\hat{\sigma}^2 = 357.94$

Exact 95% confidence interval for σ is $\left(\sqrt{\frac{5727}{\chi_{0.975,16}^2}}, \sqrt{\frac{5727}{\chi_{0.025,16}^2}} \right) = (14.090, 28.794)$

- $\hat{\sigma}_y^2 = \frac{MSC(B)-MSE}{3} = \frac{229210-358}{3} = 7628.06$

$$\hat{v} = \frac{(\hat{\sigma}_y^2)^2}{\frac{(MSC(B)/3)^2}{8} + \frac{(MSE/3)^2}{16}} = 7.975$$

A 95% confidence interval for σ_y is $\left(\sqrt{\frac{\hat{v}\hat{\sigma}_y^2}{\chi_{0.975,\hat{v}}^2}}, \sqrt{\frac{\hat{v}\hat{\sigma}_y^2}{\chi_{0.025,\hat{v}}^2}} \right) = (186.467, 529.839)$

These intervals are wider than those obtained in 3(b).

d)

- A 95% confidence interval for the difference in Time 1 and Time 2 main effects is $(\bar{y}_{1..} - \bar{y}_{2..}) \pm t_{975,16} \sqrt{\frac{2}{10} MSE} = (-98.4, -62.5)$
- A 95% confidence interval for the difference in Ad Campaign 1 and Ad Campaign 2 main effects is $(\bar{y}_{.1.} - \bar{y}_{.2.}) \pm t_{975,8} \sqrt{\frac{2}{15} MSC(B)} = (-253.4, 552.9)$. This last interval can also be obtained by doubling the limits of the interval for the fixed effect ad1 obtained when using `intervals(lmef.out)`

e) We can predict $Y_{22k} = 664.5333 + 64.2667 - 74.8667 - 0.5333 + 0 = 653.4$ for $k = new$.