

$$1. (a) E(MS_{\text{ou}}(x_u, \text{trt}))$$

$$= E \left[\frac{1}{tnm - tn} \left(\sum_{i=1}^t \sum_{j=1}^n \sum_{k=1}^m (y_{ijk} - \bar{y}_{ij\cdot})^2 \right) \right]$$

$$= \frac{1}{tnm - tn} \sum_{i=1}^t \sum_{j=1}^n \sum_{k=1}^m E (y_{ijk} - \bar{y}_{ij\cdot})^2$$

$$= \frac{1}{tn(m-1)} \sum_{i=1}^t \sum_{j=1}^n \sum_{k=1}^m E (\mu + \tau_i + u_{ij} + e_{ijk} - \mu - \tau_i - u_{ij} - \bar{e}_{ij\cdot})^2$$

$$= \frac{1}{tn(m-1)} \sum_{i=1}^t \sum_{j=1}^n \sum_{k=1}^m E (e_{ijk} - \bar{e}_{ij\cdot})^2$$

$$= \frac{1}{tn} \sum_{i=1}^t \sum_{j=1}^n E \left\{ \frac{1}{m-1} \sum_{k=1}^m (e_{ijk} - \bar{e}_{ij\cdot})^2 \right\}$$

$$= \frac{1}{tn} \sum_{i=1}^t \sum_{j=1}^n \sigma_e^2$$

$$= \frac{1}{tn} \cdot tn \sigma_e^2$$

$$= \sigma_e^2$$

1.(b). By slides 9 and 10,

②

$$E(MS_{OU}(x_u, trt))$$

$$= E\left(\frac{y'(I-P_3)y}{tn-m-tn}\right) = E\left[\frac{y'(I-P_3)}{tn(m-1)} y\right] \equiv E[y' A y],$$

where $A = \frac{I-P_3}{tn(m-1)}$

Then, $E(MS_{OU}(x_u, trt))$

$$= \text{tr}(A\Sigma) + E(y)' A E(y)$$

$$A\Sigma = \frac{1}{tn(m-1)} (I-P_3) \left[\sigma_u^2 I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m} + \sigma_e^2 I_{tnm \times tn} \right]$$

$$= \frac{1}{tn(m-1)} \left\{ \sigma_u^2 I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m} + \sigma_e^2 I_{tnm \times tn} - P_3 \left[\sigma_u^2 I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m} + \sigma_e^2 I_{tnm \times tn} \right] \right\} \quad (1)$$

Because P_3 is the projection matrix of $Z = I_{tn \times tn} \otimes \underline{\underline{1}}_{m \times 1}$, and any column of $I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m}$ is a column of Z , i.e.,

$C(I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m}) \subset C(Z)$, we have,

$$P_3 \sigma_u^2 I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m}$$

$$= \sigma_u^2 P_3 I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m}$$

$$= \sigma_u^2 I_{tn \times tn} \otimes \underline{\underline{11}}'_{m \times m} \quad (2)$$

(3)

Compute P_3 next.

$$\begin{aligned}
 P_3 &= Z(Z'Z)^{-1}Z' \\
 &= Z\left[(I_{tn \times tn} \otimes \underline{I}'_{m \times 1})(I_{tn \times tn} \otimes \underline{I}_{m \times 1})\right]^{-1}Z' \\
 &= Z\left[I_{tn \times tn} \otimes (\underline{I}'_{m \times 1} \underline{I}_{m \times 1})\right]^{-1}Z' \\
 &= Z\left[m \cdot I_{tn \times tn}\right]^{-1}Z' \\
 &= \frac{1}{m}ZZ' \\
 &= \frac{1}{m}(I_{tn \times tn} \otimes \underline{I}_{m \times 1})(I_{tn \times tn} \otimes \underline{I}'_{m \times 1}) \\
 &= \frac{1}{m}I_{tn \times tn} \otimes \underline{I}'_{m \times m}
 \end{aligned}$$

$$\therefore P_3 \sigma_e^2 I_{tn \times tn} = \sigma_e^2 P_3 = \frac{\sigma_e^2}{m} I_{tn \times tn} \otimes \underline{I}'_{m \times m} \quad - (3)$$

By (1), (2), (3),

$$A\Sigma = \frac{1}{tn(m-1)} \left\{ \sigma_e^2 I_{tn \times tn} - \frac{\sigma_e^2}{m} I_{tn \times tn} \otimes \underline{I}'_{m \times m} \right\}$$

$$\begin{aligned}
 \therefore \text{tr}(A\Sigma) &= \frac{\sigma_e^2}{tn(m-1)} \left[tn \cdot m - \frac{1}{m} \cdot tn \cdot m \right] \\
 &= \sigma_e^2 \frac{1}{tn(m-1)} \cdot tn(m-1) \\
 &= \sigma_e^2
 \end{aligned}$$

Now, consider $E(Y)'A E(Y)$.

$$\begin{aligned}
 &E(Y)'A E(Y) \\
 &= \beta'X' \frac{(I-P_3)}{tn(m-1)} X\beta = \frac{1}{tn(m-1)} \beta'X'(I-P_3)X\beta = \frac{1}{tn(m-1)} \beta'X'(X-P_3X)\beta \\
 &= 0 \quad \text{because } P_3X = X \text{ since } C(X) \subset C(P_3)
 \end{aligned}$$

$$\begin{aligned} \text{Thus, } E(MS_{\text{ou}}(x_u, \text{trt})) & \\ &= \text{tr}(A\Sigma) + E(Y)'A E(Y) \\ &= \sigma_e^2 + 0 \\ &= \sigma_e^2 \end{aligned}$$

$$1.(c) \quad \frac{MS_{\text{trt}}}{1.5\sigma_u^2 + \sigma_e^2} = \frac{(\bar{Y}_{1..} - \bar{Y}_{2..})^2}{1.5\sigma_u^2 + \sigma_e^2}, \quad \frac{MS_{\text{Xu(trt)}}}{\sigma_u^2 + \sigma_e^2} = \frac{\frac{1}{2}(Y_{111} - Y_{121})^2}{\sigma_u^2 + \sigma_e^2}$$

$$\text{Let } w_1 = \frac{\bar{Y}_{1..} - \bar{Y}_{2..}}{\sqrt{1.5\sigma_u^2 + \sigma_e^2}}, \quad w_2 = \frac{Y_{111} - Y_{121}}{\sqrt{2(\sigma_u^2 + \sigma_e^2)}}$$

$$\text{Then, } w_1^2 = \frac{MS_{\text{trt}}}{1.5\sigma_u^2 + \sigma_e^2}, \quad w_2^2 = \frac{MS_{\text{Xu(trt)}}}{\sigma_u^2 + \sigma_e^2}$$

w_1 and w_2 are normally distributed because each of them is a l.c. of normal R.V.'s.

$$\begin{aligned} \text{Consider } E(w_1) &= \frac{1}{\sqrt{1.5\sigma_u^2 + \sigma_e^2}} E(\bar{Y}_{1..} - \bar{Y}_{2..}) \\ &= \frac{1}{\sqrt{1.5\sigma_u^2 + \sigma_e^2}} (E\bar{Y}_{1..} - E\bar{Y}_{2..}) \\ &= \frac{1}{\sqrt{1.5\sigma_u^2 + \sigma_e^2}} (\mu + \tau_1 - \mu - \tau_2) \\ &= \frac{\tau_1 - \tau_2}{\sqrt{1.5\sigma_u^2 + \sigma_e^2}} \end{aligned}$$

$$\begin{aligned} \text{Var}(w_1) &= \frac{1}{1.5\sigma_u^2 + \sigma_e^2} \text{Var}(\bar{Y}_{1..} - \bar{Y}_{2..}) \\ &= \frac{1}{1.5\sigma_u^2 + \sigma_e^2} \text{Var}\left[\frac{1}{2}(Y_{111} + Y_{121}) - \frac{1}{2}(Y_{211} + Y_{212})\right] \\ &= \frac{1}{1.5\sigma_u^2 + \sigma_e^2} \cdot \frac{1}{4} \left[\text{Var}(Y_{111}) + \text{Var}(Y_{121}) + \text{Var}(Y_{211}) + \text{Var}(Y_{212}) + 2\text{Cov}(Y_{211}, Y_{212}) \right] \\ &= \frac{1}{1.5\sigma_u^2 + \sigma_e^2} \cdot \frac{1}{4} \left[4(\sigma_u^2 + \sigma_e^2) + 2\sigma_u^2 \right] \\ &= 1 \end{aligned}$$

$$\therefore w_1 \sim N\left(\frac{\tau_1 - \tau_2}{\sqrt{1.5\sigma_u^2 + \sigma_e^2}}, 1\right)$$

$$\therefore \frac{MS_{trt}}{1.5\sigma_u^2 + \sigma_e^2} \sim \chi_1^2 \left(\frac{(\tau_1 - \tau_2)^2}{1.5\sigma_u^2 + \sigma_e^2} \right) \quad (1)$$

Similarly, $E(W_2) = \frac{1}{\sqrt{2(\sigma_u^2 + \sigma_e^2)}} [E(Y_{111}) - E(Y_{121})]$

$$= \frac{1}{\sqrt{2(\sigma_u^2 + \sigma_e^2)}} (\mu + \tau_1 - \mu - \tau_1)$$

$$= 0$$

$$\text{Var}(W_2) = \frac{1}{2(\sigma_u^2 + \sigma_e^2)} \text{Var}(Y_{111} - Y_{121})$$

$$= \frac{1}{2(\sigma_u^2 + \sigma_e^2)} \left[\text{Var}(Y_{111}) + \text{Var}(Y_{121}) - 2\text{Cov}(Y_{111}, Y_{121}) \right]$$

$$= \frac{1}{2(\sigma_u^2 + \sigma_e^2)} [2(\sigma_u^2 + \sigma_e^2) - 0]$$

$$= 1$$

$\therefore W_2 \sim N(0, 1)$

$$\therefore \frac{MS_{Xu(trt)}}{\sigma_u^2 + \sigma_e^2} \sim \chi_1^2(0) \quad (2)$$

$$\text{Cov}(W_1, W_2) = \text{Cov} \left(\frac{\bar{Y}_{1..} - \bar{Y}_{2..}}{\sqrt{1.5\sigma_u^2 + \sigma_e^2}}, \frac{Y_{111} - Y_{121}}{\sqrt{2(\sigma_u^2 + \sigma_e^2)}} \right)$$

$$= \frac{1}{\sqrt{(1.5\sigma_u^2 + \sigma_e^2)} \sqrt{2(\sigma_u^2 + \sigma_e^2)}} \text{Cov} \left(\frac{1}{2}(Y_{111} + Y_{121}) - \frac{1}{2}(Y_{211} + Y_{212}), Y_{111} - Y_{121} \right)$$

$$= \frac{1}{\sqrt{1.5\sigma_u^2 + \sigma_e^2} \sqrt{2(\sigma_u^2 + \sigma_e^2)}} \left\{ \text{Var}(Y_{111}) - \text{Var}(Y_{121}) \right\} / 2$$

$$= \frac{1}{\sqrt{1.5\sigma_u^2 + \sigma_e^2} \sqrt{2(\sigma_u^2 + \sigma_e^2)}} (\sigma_u^2 + \sigma_e^2 - \sigma_u^2 - \sigma_e^2) / 2$$

$$= 0$$

$\therefore W_1, W_2$ are independent

$\therefore \frac{MS_{trt}}{1.5\sigma_u^2 + \sigma_e^2}, \frac{MS_{xu(trt)}}{\sigma_u^2 + \sigma_e^2}$ are independent (3)

By (1), (2), (3),

$$\frac{MS_{trt}}{1.5\sigma_u^2 + \sigma_e^2} / \frac{MS_{xu(trt)}}{\sigma_u^2 + \sigma_e^2} \sim F_{1,1} \left(\frac{(\tau_1 - \tau_2)^2}{1.5\sigma_u^2 + \sigma_e^2} \right)$$

2.(a) Source	D.f.	Sum of Square
Diet	$D-1$	$\sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{i..} - \bar{y}_{...})^2$
Drug	$M-1$	$\sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{.j.} - \bar{y}_{...})^2$
Diet x Drug	$(D-1)(M-1)$	$\sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{ij.} - \bar{y}_{i..} - \bar{y}_{.j.} + \bar{y}_{...})^2$
Litter (Diet)	$D(L-1)$	$\sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{i.k} - \bar{y}_{i..})^2$
Error	$DML - DM - DL + D$	$\sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (y_{ijk} - \bar{y}_{ij.} - \bar{y}_{i.k} + \bar{y}_{i..})^2$
Corrected Total	$DML - 1$	$\sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (y_{ijk} - \bar{y}_{...})^2$

The above notations are based on the model:

$$y_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + u_{ik} + e_{ijk},$$

where $i=1, 2, \dots, D$; $j=1, 2, \dots, M$; $k=1, 2, \dots, L$;

u_{ik} is the random effect of litters, $u_{ik} \text{ iid } N(0, \sigma_u^2)$;

$e_{ijk} \text{ iid } N(0, \sigma_e^2)$.

Note that "Error" corresponds to Drug x Litter (Diet), which

has $(M-1)D(L-1) = DML - DM - DL + D$ degrees of freedom.

2.(b) $E(MS_{Diet})$

$$\begin{aligned}
&= \frac{1}{D-1} E \left\{ \sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{i..} - \bar{y}_{...})^2 \right\} \\
&= \frac{1}{D-1} \cdot ML \cdot \sum_{i=1}^D E(\bar{y}_{i..} - \bar{y}_{...})^2 \\
&= \frac{ML}{D-1} \sum_{i=1}^D E \left[\mu + \alpha_i + \beta_{.} + \delta_{i.} + \bar{u}_{i.} + \bar{e}_{i..} - \mu - \bar{\alpha}_{.} - \bar{\beta}_{.} - \bar{\gamma}_{..} - \bar{u}_{..} - \bar{e}_{...} \right]^2 \\
&= \frac{ML}{D-1} \left\{ \sum_{i=1}^D (\alpha_i - \bar{\alpha}_{.} + \bar{\gamma}_{i.} - \bar{\gamma}_{..})^2 + \sum_{i=1}^D E(\bar{u}_{i.} - \bar{u}_{..})^2 + \sum_{i=1}^D E(\bar{e}_{i..} - \bar{e}_{...})^2 \right\} \\
&= \frac{ML}{D-1} \left[\sum_{i=1}^D (\alpha_i - \bar{\alpha}_{.} + \bar{\gamma}_{i.} - \bar{\gamma}_{..})^2 \right] + ML E \left[\frac{\sum_{i=1}^D (\bar{u}_{i.} - \bar{u}_{..})^2}{D-1} \right] + ML E \left[\frac{\sum_{i=1}^D (\bar{e}_{i..} - \bar{e}_{...})^2}{D-1} \right] \\
&= \frac{ML}{D-1} \left[\sum_{i=1}^D (\alpha_i - \bar{\alpha}_{.} + \bar{\gamma}_{i.} - \bar{\gamma}_{..})^2 \right] + ML \text{Var}(\bar{u}_{i.}) + ML \cdot \text{Var}(\bar{e}_{i..}) \\
&= \frac{ML}{D-1} \left[\sum_{i=1}^D (\alpha_i - \bar{\alpha}_{.} + \bar{\gamma}_{i.} - \bar{\gamma}_{..})^2 \right] + ML \cdot \frac{\sigma_u^2}{L} + ML \cdot \frac{\sigma_e^2}{ML} \\
&= \frac{ML}{D-1} \left[\sum_{i=1}^D (\alpha_i - \bar{\alpha}_{.} + \bar{\gamma}_{i.} - \bar{\gamma}_{..})^2 \right] + M\sigma_u^2 + \sigma_e^2
\end{aligned}$$

$E(MS_{Drug})$

$$\begin{aligned}
&= \frac{1}{M-1} E \left\{ \sum_{j=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{.j.} - \bar{y}_{...})^2 \right\} \\
&= \frac{DL}{M-1} \sum_{j=1}^M E(\bar{y}_{.j.} - \bar{y}_{...})^2 \\
&= \frac{DL}{M-1} \sum_{j=1}^M E(\mu + \bar{\alpha}_{.} + \beta_j + \delta_{.j} + \bar{u}_{.j.} + \bar{e}_{.j.} - \mu - \bar{\alpha}_{.} - \bar{\beta}_{.} - \bar{\gamma}_{..} - \bar{u}_{..} - \bar{e}_{...})^2 \\
&= \frac{DL}{M-1} \sum_{j=1}^M E(\beta_j - \bar{\beta}_{.} + \bar{\delta}_{.j} - \bar{\gamma}_{..} + \bar{e}_{.j.} - \bar{e}_{...})^2
\end{aligned}$$

$$= \frac{DL}{M-1} \sum_{j=1}^M (\beta_j - \bar{\beta} + \bar{\gamma}_{.j} - \bar{\gamma}_{..})^2 + \frac{DL}{M-1} E \left\{ \sum_{j=1}^M (\bar{e}_{.j} - \bar{e}_{...})^2 \right\}$$

$$= \frac{DL}{M-1} \sum_{j=1}^M (\beta_j - \bar{\beta} + \bar{\gamma}_{.j} - \bar{\gamma}_{..})^2 + DL \text{Var}(\bar{e}_{.j})$$

$$= \frac{DL}{M-1} \sum_{j=1}^M (\beta_j - \bar{\beta} + \bar{\gamma}_{.j} - \bar{\gamma}_{..})^2 + DL \cdot \frac{\sigma_e^2}{DL}$$

$$= \frac{DL}{M-1} \sum_{j=1}^M (\beta_j - \bar{\beta} + \bar{\gamma}_{.j} - \bar{\gamma}_{..})^2 + \sigma_e^2$$

$E(MS_{\text{Diet} \times \text{Drug}})$

$$= \frac{1}{(D-1)(M-1)} E \left\{ \sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{ij.} - \bar{y}_{i..} - \bar{y}_{.j.} + \bar{y}_{...})^2 \right\}$$

$$= \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M E(\bar{y}_{ij.} - \bar{y}_{i..} - \bar{y}_{.j.} + \bar{y}_{...})^2$$

$$= \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M E(\alpha_i + \beta_j + \gamma_{ij} + \mu_{i.} + \bar{e}_{ij.} - \alpha_i - \bar{\beta} - \bar{\gamma}_{i.} - \mu_{i.} - \bar{e}_{i..} - \bar{\alpha} - \beta_j - \bar{\gamma}_{.j} - \mu_{..} - \bar{e}_{.j.} + \bar{\alpha} + \bar{\beta} + \bar{\gamma}_{..} + \mu_{..} + \bar{e}_{...})^2$$

$$= \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M E(\gamma_{ij} - \bar{\gamma}_{i.} - \bar{\gamma}_{.j} + \bar{\gamma}_{..} + \bar{e}_{ij.} - \bar{e}_{i..} - \bar{e}_{.j.} + \bar{e}_{...})^2$$

$$= \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M (\gamma_{ij} - \bar{\gamma}_{i.} - \bar{\gamma}_{.j} + \bar{\gamma}_{..})^2 + LE \left\{ \frac{1}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M (\bar{e}_{ij.} - \bar{e}_{i..} - \bar{e}_{.j.} + \bar{e}_{...})^2 \right\}$$

$$= \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M (\gamma_{ij} - \bar{\gamma}_{i.} - \bar{\gamma}_{.j} + \bar{\gamma}_{..})^2 + \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M \text{Var}(\bar{e}_{ij.} - \bar{e}_{i..} - \bar{e}_{.j.} + \bar{e}_{...})$$

--- (A)

where $\text{Var}(\bar{e}_{ij} = \bar{e}_{i\cdot} - \bar{e}_{\cdot j} + \bar{e}_{\dots})$

$$\begin{aligned}
 &= \text{Var}(\bar{e}_{ij}) + \text{Var}(\bar{e}_{i\cdot}) + \text{Var}(\bar{e}_{\cdot j}) + \text{Var}(\bar{e}_{\dots}) - 2\text{Cov}(\bar{e}_{ij}, \bar{e}_{i\cdot}) \\
 &\quad - 2\text{Cov}(\bar{e}_{ij}, \bar{e}_{\cdot j}) + 2\text{Cov}(\bar{e}_{ij}, \bar{e}_{\dots}) - 2\text{Cov}(\bar{e}_{i\cdot}, \bar{e}_{\dots}) \\
 &\quad - 2\text{Cov}(\bar{e}_{\cdot j}, \bar{e}_{\dots}) + 2\text{Cov}(\bar{e}_{i\cdot}, \bar{e}_{\cdot j}) \\
 &= \frac{\sigma_e^2}{L} + \frac{\sigma_e^2}{ML} + \frac{\sigma_e^2}{DL} + \frac{\sigma_e^2}{DML} - 2 \frac{1}{M} \frac{\sigma_e^2}{L} - \frac{2\sigma_e^2}{D \cdot L} + \frac{2\sigma_e^2}{DML} \\
 &\quad - \frac{2\sigma_e^2}{DML} - \frac{2\sigma_e^2}{DML} + \frac{2\sigma_e^2}{DML} \\
 &= \sigma_e^2 \left(\frac{1}{L} - \frac{1}{ML} - \frac{1}{DL} + \frac{1}{DML} \right) \\
 &= \sigma_e^2 (D-1)(M-1) / DML
 \end{aligned}$$

By \textcircled{A} , $E(MS_{\text{Diet} \times \text{Drug}})$

$$\begin{aligned}
 &= \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M (\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}_{\dots})^2 + \frac{LDM}{(D-1)(M-1)} \frac{\sigma_e^2 (D-1)(M-1)}{DML} \\
 &= \frac{L}{(D-1)(M-1)} \sum_{i=1}^D \sum_{j=1}^M (\bar{y}_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}_{\dots})^2 + \sigma_e^2
 \end{aligned}$$

$E(MS_{\text{Litter(Diet)}})$

$$\begin{aligned}
 &= \frac{1}{D(L-1)} E \left\{ \sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (\bar{y}_{i\cdot k} - \bar{y}_{i\cdot})^2 \right\} + \dots \\
 &= \frac{1}{D(L-1)} M \sum_{i=1}^D \sum_{k=1}^L E(\bar{y}_{i\cdot k} - \bar{y}_{i\cdot})^2 \\
 &= \frac{M}{D(L-1)} \sum_{i=1}^D \sum_{k=1}^L E(u_{ik} - \bar{u}_{i\cdot} + \bar{e}_{i\cdot k} - e_{i\cdot})^2
 \end{aligned}$$

$$= \frac{M}{D} \sum_{i=1}^D \left\{ E \left\{ \frac{\sum_{k=1}^L (u_{ik} - \bar{u}_i)^2}{L-1} \right\} + E \left\{ \frac{\sum_{k=1}^L (e_{ik} - \bar{e}_{i..})^2}{L-1} \right\} \right\}$$

$$= \frac{M}{D} \sum_{i=1}^D \left\{ \sigma_u^2 + \frac{\sigma_e^2}{M} \right\}$$

$$= M\sigma_u^2 + \sigma_e^2$$

$E(MSE_{Error})$

$$= \frac{1}{DML - DM - DL + D} \cdot E \left\{ \sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L (y_{ijk} - \bar{y}_{ij.} - \bar{y}_{i.k} + \bar{y}_{i..})^2 \right\}$$

$$= \frac{1}{D(M-1)(L-1)} \sum_{i=1}^D \sum_{j=1}^M \sum_{k=1}^L E (e_{ijk} - \bar{e}_{ij.} - \bar{e}_{i.k} + \bar{e}_{i..})^2$$

$$= \frac{1}{D(M-1)(L-1)} \cdot \sum_{i=1}^D E \left\{ \sum_{j=1}^M \sum_{k=1}^L (e_{ijk} - \bar{e}_{ij.} - \bar{e}_{i.k} + \bar{e}_{i..})^2 \right\}$$

$$= \frac{1}{D(M-1)(L-1)} \sum_{i=1}^D [(M-1)(L-1)\sigma_e^2]$$

$$= \sigma_e^2$$

$$Z(c) \frac{SS_{\text{Litter(Diet)}}}{M\sigma_u^2 + \sigma_e^2} \sim \chi^2_{D(L-1)}$$

$$\frac{SS_{\text{Error}}}{\sigma_e^2} \sim \chi^2_{D(M-1)(L-1)}$$

2. (d) To test for diet main effects, we can use

$$F = \frac{MS_{\text{Diet}}}{MS_{\text{Litter(Diet)}}},$$

which has a central F dist'n with $(D-1, D(L-1))$ d.f.

under H_0 .

2(e) To test for drug main effects, we can use

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$$F = \frac{MS_{\text{Drug}}}{MS_{\text{Error}}}$$

which has a central F dist'n with $(M-1, D(M-1)(L-1))$ d.f.

under H_0 .

2.(f) To test for diet by drug interactions, we can use

$$F = \frac{MS_{Diet \times Drug}}{MS_{Error}}$$

which has a central F dist'n with $(D-1)(M-1), D(M-1)(L-1)$

d.f. under H_0 .

2(g) The difference between the drug 1 mean and the drug 2 mean when averaging over diets is:

$$\beta_1 + \bar{\gamma}_{.1} - \beta_2 - \bar{\gamma}_{.2}$$

Consider $\bar{y}_{.1} - \bar{y}_{.2}$.

$$\begin{aligned} \bar{y}_{.1} - \bar{y}_{.2} &= (\mu + \bar{\alpha}_{.} + \beta_1 + \bar{\gamma}_{.1} + \bar{u}_{..} + \bar{e}_{.1}) - (\mu + \bar{\alpha}_{.} + \beta_2 + \bar{\gamma}_{.2} + \bar{u}_{..} + \bar{e}_{.2}) \\ &= (\beta_1 + \bar{\gamma}_{.1} - \beta_2 - \bar{\gamma}_{.2}) + (\bar{e}_{.1} - \bar{e}_{.2}) \end{aligned}$$

$$\therefore \bar{y}_{.1} - \bar{y}_{.2} \sim N(\beta_1 + \bar{\gamma}_{.1} - \beta_2 - \bar{\gamma}_{.2}, \frac{2\sigma_e^2}{DL})$$

$$\therefore E(MS_{Error}) = \sigma_e^2 \quad \text{and} \quad \frac{MS_{Error} \cdot D(M-1)(L-1)}{\sigma_e^2} \sim \chi_{D(M-1)(L-1)}^2 \quad \text{by (c)}$$

$$\therefore \frac{(\bar{y}_{.1} - \bar{y}_{.2}) - (\beta_1 + \bar{\gamma}_{.1} - \beta_2 - \bar{\gamma}_{.2})}{\sqrt{\frac{2MS_{Error}}{DL}}} \sim t_{D(M-1)(L-1)}$$

\therefore A 95% confidence interval for the desired difference is

$$(\bar{y}_{.1} - \bar{y}_{.2} - t_{D(M-1)(L-1), 0.975} \sqrt{\frac{2MS_{Error}}{DL}}, \bar{y}_{.1} - \bar{y}_{.2} + t_{D(M-1)(L-1), 0.975} \sqrt{\frac{2MS_{Error}}{DL}})$$

2(h) The difference between the diet 1 drug 1 mean and the diet 2 drug 1 mean is:

$$\alpha_1 + \gamma_{11} - \alpha_2 - \gamma_{21}$$

Consider $\bar{y}_{11.} - \bar{y}_{21.}$.

$$\begin{aligned} & \bar{y}_{11.} - \bar{y}_{21.} \\ &= (\mu + \alpha_1 + \beta_1 + \gamma_{11} + \bar{u}_{1.} + \bar{e}_{11.}) - (\mu + \alpha_2 + \beta_1 + \gamma_{21} + \bar{u}_{2.} + \bar{e}_{21.}) \\ &= (\alpha_1 + \gamma_{11} - \alpha_2 - \gamma_{21}) + (\bar{u}_{1.} + \bar{e}_{11.} - \bar{u}_{2.} - \bar{e}_{21.}) \end{aligned}$$

$$\begin{aligned} \text{Var}(\bar{y}_{11.} - \bar{y}_{21.}) &= \text{Var}(\bar{u}_{1.} - \bar{u}_{2.} + \bar{e}_{11.} - \bar{e}_{21.}) \\ &= \text{Var}(\bar{u}_{1.} - \bar{u}_{2.}) + \text{Var}(\bar{e}_{11.} - \bar{e}_{21.}) \\ &= \frac{2\sigma_u^2}{L} + \frac{2\sigma_e^2}{L} \end{aligned}$$

$$\therefore \bar{y}_{11.} - \bar{y}_{21.} \sim N\left(\alpha_1 + \gamma_{11} - \alpha_2 - \gamma_{21}, \frac{2(\sigma_u^2 + \sigma_e^2)}{L}\right)$$

Because $E(MS_{\text{Litter(Diet)}}) = M\sigma_u^2 + \sigma_e^2$ and $E(MS_{\text{Error}}) = \sigma_e^2$,

we have $\hat{\sigma}_e^2 = MS_{\text{Error}}$ and $\hat{\sigma}_u^2 = \frac{1}{M} (MS_{\text{Litter(Diet)}} - MS_{\text{Error}})$. (*)

Then, by Cochran-Satterthwaite approximation,

$$d. \left[\frac{2(\hat{\sigma}_u^2 + \hat{\sigma}_e^2)}{L} / \frac{2(\sigma_u^2 + \sigma_e^2)}{L} \right] \sim \chi_d^2$$

$$\text{where } d = \frac{\left\{ \frac{2}{L} \left[\frac{1}{M} (MS_{\text{Litter(Diet)}} - MS_{\text{Error}}) + MS_{\text{Error}} \right] \right\}^2}{\frac{\left[\frac{2}{L} \cdot \frac{1}{M} (MS_{\text{Litter(Diet)}}) \right]^2}{D(L-1)} + \frac{\left[\frac{2}{L} \left(1 - \frac{1}{M}\right) MS_{\text{Error}} \right]^2}{D(M-1)(L-1)}}$$

Thus, $\frac{\bar{Y}_{11} - \bar{Y}_{21} - (\alpha_1 + \delta_{11} - \alpha_2 - \delta_{21})}{\sqrt{\frac{2(\hat{\sigma}_u^2 + \hat{\sigma}_e^2)}{L}}} \sim t_d$

∴ a 95% confidence interval for the desired difference is

$$(\bar{Y}_{11} - \bar{Y}_{21} - t_{d,0.975} \sqrt{\frac{2(\hat{\sigma}_u^2 + \hat{\sigma}_e^2)}{L}}, \bar{Y}_{11} - \bar{Y}_{21} + t_{d,0.975} \sqrt{\frac{2(\hat{\sigma}_u^2 + \hat{\sigma}_e^2)}{L}})$$

where $\hat{\sigma}_u^2$ and $\hat{\sigma}_e^2$ are as in (*)

2(i) An average response for each litter is

$$\bar{y}_{i \cdot k} = \mu + \alpha_i + \bar{\beta}_0 + \bar{\gamma}_{ii} + u_{ik} + \bar{e}_{ik},$$

for $i=1, \dots, D; k=1, \dots, L$.

2(j) Let $Z_{ik} = \bar{Y}_{i \cdot k}$, $\tau_i = \mu + \alpha_i + \bar{\beta} + \bar{\gamma}_i$, $\epsilon_{ik} = u_{ik} + \bar{e}_{i \cdot k}$, (21)

then, $\epsilon_{ik} \sim N(0, \sigma_u^2 + \frac{\sigma_e^2}{M})$.

By part (i),

$$Z_{ik} = \tau_i + \epsilon_{ik}, \quad \text{for } i=1, \dots, D; k=1, \dots, L.$$

This model has the form of ^{the usual model for} a completely randomized design.

The analysis using this model can give us the inference about the main effects of diets.

The results of such an analysis would be the same with those obtained using the full data set.

