

## Stat 543 Assignment 11

### Asymptotics of the LRT and Posterior Distributions, and Wald and Score Tests

1. (More Estimation in a Zero-Inflated Poisson Model) Consider the situation of Problem 1 of Assignment 10, and in particular, inference based on the  $n = 20$  observations Vardeman simulated from the distribution.

- (a) Find a large sample 90% joint confidence region for  $(p, \lambda)$  based on the loglikelihood function (based on inverting LRT's). Plot this in the  $(p, \lambda)$ -plane and to the extent possible, compare it to the elliptical region you found in Assignment 10.
- (b) Note that for a fixed value of  $\lambda$ ,

$$p = \frac{\frac{n_0}{n} - 1}{\exp(-\lambda) - 1}$$

maximizes the likelihood. Use this fact to find and plot the profile loglikelihood for  $\lambda$ . Use this plot and make an approximate 90% confidence interval for  $\lambda$ . How does this interval compare to the one you found in part e) of Problem 1 from Assignment 10?

2. Consider again the model of Problem 2 of Assignment 10. Below are  $n = 20$  observations that Vardeman simulated from this model.

1.36, 1.35, 0.78, 1.85, 2.32, 0.55, 1.07, -0.57, -0.38, 0.25,  
-0.36, 1.71, 1.40, 0.46, 3.16, -0.78, 0.69, -0.03, 1.26, 0.44

- (a) Plot the loglikelihood for this sample. What, approximately, is the maximum likelihood estimate for  $\alpha$ ?
  - (b) If you wished to test the hypothesis  $H_0: \alpha = .4$  with Type I error probability .1, what would be your decision here? Carefully explain. (Use a likelihood ratio test).
  - (c) Use the plot from a) and make an approximate 90% confidence interval for  $\alpha$  based on the likelihood function (based on inverting LRT's). Use the method of f) of Problem 2 on Assignment 10 and make another approximate 90% interval. How do these 2 intervals compare?
3. Problems 1.2.4 and 1.2.5 of B&D.
  4. Consider Bayesian inference for the binomial parameter  $p$ . In particular, for sake of convenience, consider the Uniform  $(0, 1)$  (Beta( $\alpha, \beta$ ) for  $\alpha = \beta = 1$ ) prior distribution.
    - (a) It is possible to argue from reasonably elementary principles that in this binomial context, where  $\Theta = (0, 1)$ , the Beta posteriors have a consistency property. That

is, simple arguments can be used to show that for any fixed  $p_0$  and any  $\epsilon > 0$ , for  $X_n \sim \text{binomial}(n, p_0)$ , the random variable

$$Y_n = \int_{p_0-\epsilon}^{p_0+\epsilon} \frac{1}{B(\alpha + X_n, \beta + (n - X_n))} p^{\alpha+X_n-1} (1-p)^{\beta+(n-X_n)-1} dp$$

(which is the posterior probability assigned to the interval  $(p_0 - \epsilon, p_0 + \epsilon)$ ) converges in  $p_0$  probability to 1 as  $n \rightarrow \infty$ . This part of the problem is meant to lead you through this argument. Let  $\epsilon > 0$  and  $\delta > 0$ .

i) Argue that there exists  $m$  such that if  $n \geq m$ ,  $\left| \frac{x_n}{n} - \frac{\alpha+x_n}{\alpha+\beta+n} \right| < \frac{\epsilon}{3} \forall x_n = 0, 1, \dots, n$ .

ii) Note that the posterior variance is  $\frac{(\alpha+x_n)(\beta+n-x_n)}{(\alpha+\beta+n)^2(\alpha+\beta+n+1)}$ . Argue there is an  $m'$  such that if  $n \geq m'$  the probability that the posterior assigns to  $\left( \frac{\alpha+x_n}{\alpha+\beta+n} - \frac{\epsilon}{3}, \frac{\alpha+x_n}{\alpha+\beta+n} + \frac{\epsilon}{3} \right)$  is at least  $1 - \delta \forall x_n = 0, 1, \dots, n$ .

iii) Argue that there is an  $m''$  such that if  $n \geq m''$  the  $p_0$  probability that  $\left| \frac{X_n}{n} - p_0 \right| < \frac{\epsilon}{3}$  is at least  $1 - \delta$ .

Then note that if  $n \geq \max(m, m', m'')$  i) and ii) together imply that the posterior probability assigned to  $\left( \frac{x_n}{n} - \frac{2\epsilon}{3}, \frac{x_n}{n} + \frac{2\epsilon}{3} \right)$  is at least  $1 - \delta$  for any realization  $x_n$ . Then provided  $\left| \frac{x_n}{n} - p_0 \right| < \frac{\epsilon}{3}$  the posterior probability assigned to  $(p_0 - \epsilon, p_0 + \epsilon)$  is also at least  $1 - \delta$ . But iii) says this happens with  $p_0$  probability at least  $1 - \delta$ . That is, for large  $n$ , with  $p_0$  probability at least  $1 - \delta$ ,  $Y_n \geq 1 - \delta$ . Since  $\delta$  is arbitrary, (and  $Y_n \leq 1$ ) we have the convergence of  $Y_n$  to 1 in  $p_0$  probability.

(b) Vardeman intends to argue in class that posterior densities for large  $n$  tend to look normal (with means and variances related to the likelihood material). The posteriors in this binomial problem are Beta  $(\alpha + x_n, \beta + (n - x_n))$  (and we can think of  $X_n \sim \text{Bi}(n, p_0)$  as derived as the sum of  $n$  iid Bernoulli  $(p_0)$  variables). So we ought to expect Beta distributions for large parameter values to look roughly normal. To illustrate this do the following. For  $\rho = .3$  (for example ... any other value would do as well), consider the Beta  $(\alpha + n\rho, \beta + n(1 - \rho))$  (posterior) distributions for  $n = 10, 20, 40$  and  $100$ . For  $p_n \sim \text{Beta}(\alpha + n\rho, \beta + n(1 - \rho))$  plot the probability densities for the variables

$$\sqrt{\frac{n}{\rho(1-\rho)}} (p_n - \rho)$$

on a single set of axes along with the standard normal density. Note that if  $W$  has pdf  $f(\cdot)$ , then  $aW + b$  has pdf  $g(\cdot) = \frac{1}{a} f\left(\frac{\cdot-b}{a}\right)$ . (Your plots are translated and rescaled posterior densities of  $p$  based on possible observed values  $x_n = .3n$ .) If this is any help in doing this plotting, Vardeman tried to calculate values of the Beta function using MathCad and got the following:  $(B(4, 8))^{-1} = 1.32 \times 10^3$ ,  $(B(7, 15))^{-1} = 8.14 \times 10^5$ ,  $(B(13, 29))^{-1} = 2.291 \times 10^{11}$  and  $(B(31, 71))^{-1} = 2.967 \times 10^{27}$ .

5. Suppose that  $X, Y$  and  $Z$  are independent binomial variables,  $X \sim \text{bin}(n, p_1)$ ,  $Y \sim \text{bin}(n, p_2)$  and  $Z \sim \text{bin}(n, p_3)$ . For the parameter space (for  $(p_1, p_2, p_3)$ )  $\Theta = [0, 1]^3$ , we will consider testing  $H_0: p_1 = p_2 = p_3$  based on  $(X, Y, Z)$ .

- (a) Find the general forms of the likelihood ratio tests, the Wald tests and the score tests of this hypothesis.
- (b) Use the fact that the parameter space here is basically 3-dimensional while  $\Theta_0$  is basically 1-dimensional so that there are 2 independent constraints involved and the limiting  $\chi^2$  distributions of the test statistics thus have  $\nu = 2$  associated degrees of freedom to actually carry out these tests with  $\alpha \approx .05$  if  $X = 33$ ,  $Y = 53$  and  $Z = 59$ , all based on  $n = 100$ .